

Univariate Side Channel Attacks and Leakage Modeling

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- 1 Introduction
 - Problematics
 - Motivations
- 2 Description
 - Preliminaries
 - Description
- 3 Our Work - Analyses
 - All DPAs reduce to PPA
 - From PPA to CPA
 - Model-Free
 - Experiments
- 4 Conclusion



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4 Conclusion



- Focus on **Univariate First-Order** SCA
 - ▶ SCA focusses on a single intermediate variable manipulation.
 - ▶ we don't deal with higher-order SCA!
- What do we want to do?
 - ▶ Exhibit a linear dependency
- Which tools do we have in the literature?
 - ▶ DPA



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- What well-established theory says about dependency test?
 - ▶ Linear correlation: Pearson's coefficient
 - ▶ Quantity of shared Information: mutual information (or Kullback-Leibler Distance)



It's round, it rolls ... it's a wheel !

Linear correlation

- Measures linear dependencies between two sets
- Observations and estimations
- Estimations ?



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Linear estimation

- Exhibits a linear model from observations
- What to do with this model ?



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- Study the relationship between some well-used distinguishers: DPA and variants, PPA and CPA.
 - ▶ Do they differ one to each other?
 - ▶ Do they involve different statistical tools?
 - ▶ Is one of them always better than the other ones?
 - Study some distinguishers that do not need an *a priori* modeling of the observations.
- ⇒ We rule on how to lead a linear univariate side-channel attack!



- (Noisy) Observation of a processing:

$$L \longleftarrow V_k = S(k, P)$$

- $(p_i)_i$: known values taken by P .
- $(v_{k,i})_i$: **unknown** values taken by V_k .
- $(\ell_{k,i})_i$: known values taken by L .

Assumption on Leverages

- Composed of two parts:
 - ▶ A deterministic part $\delta(\cdot)$,
 - ▶ An **independent** noise B ,
- such that $L = \delta(V_k) + B$.

1. **Measurement** : get a sample $(\ell_{k,i})_i$ related to a sample $(p_i)_i$ of plaintexts.
2. **Model Selection** : Design/Select a function $\mathbf{m}(\cdot)$ to model $\delta(\cdot)$.
3. **Prediction** : For every \hat{k} , compute

$$m_{\hat{k},i} = \mathbf{m}(v_{\hat{k},i}) = \mathbf{m}(S(\hat{k}, p_i)) .$$

The r.v. corresponding to $m_{\hat{k},i}$ is denoted by $M_{\hat{k}}$.

4. **Distinguisher Selection** : Choose a statistical distinguisher Δ .
5. **Key Discrimination** : For every \hat{k} , compute the **distinguishing value** $\Delta_{\hat{k}}$:

$$\Delta_{\hat{k}} = \Delta \left((\ell_{k,i})_i, (m_{\hat{k},i})_i \right) .$$

6. **Key Candidate Selection** : Output the o key hypotheses that maximize $\Delta_{\hat{k}}$.

What differs from an attack to another

- The distinguisher Δ
- The model m

→ How to choose the distinguisher?

→ How to choose the model?

⇒ Link between the two questions?



single-bit DPA

- Example: $\mathbf{m}(v_{\hat{k},i}) = S_0(p_i \oplus \hat{k})$, where S_0 is the 0th coordinate function of S .
- $\text{Im}(\mathbf{m}) = \{w_0, w_1\}$
- (e.g. $w_0 = 0$ and $w_1 = 1$).
- $\Delta_{\hat{k}} = \hat{\mathbb{E}}(L \mid M_{\hat{k}} = w_0) - \hat{\mathbb{E}}(L \mid M_{\hat{k}} = w_1) = \text{SB-DPA}(\hat{k})$.



Differential Power Analysis

single-bit DPA

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all-or-nothing DPA

- Example: $\mathbf{m}(v_{\hat{k},i}) = \text{HW}(S(p_i \oplus \hat{k}))$.
- $\text{Im}(\mathbf{m}) = \{\dots, \omega_0, \dots, \omega_1, \dots\}$
- (e.g. $\text{Im}(\mathbf{m}) = \{0, \dots, 8\}$ and $w_0 = 0$ and $w_1 = 8$).
- $\Delta_{\hat{k}} = \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_0) - \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_1) = \text{AON-DPA}(\hat{k})$.



generalized DPA

- Example: $m(v_{\hat{k},i}) = \text{HW}(S(p_i \oplus \hat{k}))$.
- $\text{Im}(m) = \dots \cup \Omega_0 \cup \dots \cup \Omega_1 \cup \dots$
- (e.g. $\text{Im}(m) = \Omega_0 \cup \Omega_1$ and $\Omega_0 = \{0, \dots, 4\}$ and $\Omega_1 = \{5, \dots, 8\}$).
- $\Delta_{\hat{k}} = \hat{\mathbb{E}}(L \mid M_{\hat{k}} \in \Omega_0) - \hat{\mathbb{E}}(L \mid M_{\hat{k}} \in \Omega_1) = \mathbf{G-DPA}(\hat{k})$.

PPA

- Example: $\mathbf{m}(v_{\hat{k},i}) = \text{HW}(S(p_i \oplus \hat{k}))$.
- $\text{Im}(\mathbf{m}) = \{\omega_i; i\}$ (e.g. $\text{Im}(\mathbf{m}) = \{0, \dots, 8\}$).
- $\Delta_{\hat{k}} = \sum_{\omega_i \in \text{Im}(\mathbf{m})} \alpha_i \cdot \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_i) = \text{PPA}_{(\alpha_i)_i}(\hat{k})$,
- $\alpha_i \in \mathbb{R}$

CPA

- Example: $\mathbf{m}(v_{\hat{k},i}) = \text{HW}(S(p_i \oplus \hat{k}))$.
- $\text{Im}(\mathbf{m}) = \{\omega_i; i\}$
- (e.g. $\text{Im}(\mathbf{m}) = \{0, \dots, 8\}$).
- $\Delta_{\hat{k}} = \hat{\rho}(L, M_{\hat{k}}) = \frac{\widehat{\text{cov}}(L, M_{\hat{k}})}{\hat{\sigma}(L) \cdot \hat{\sigma}(M_{\hat{k}})} = \text{CPA}(\hat{k})$,
- $\hat{\sigma}(L)$ and $\hat{\sigma}(M_{\hat{k}})$: estimators of the standard dev. of L and $V_{\hat{k}}$.
- $\widehat{\text{cov}}(L, M_{\hat{k}})$: estimator of the covariance between L and $V_{\hat{k}}$
- Note: $\widehat{\text{cov}}(L, M_{\hat{k}}) = \widehat{\mathbb{E}}(LM_{\hat{k}}) - \widehat{\mathbb{E}}(L)\widehat{\mathbb{E}}(M_{\hat{k}})$.

Target Uniformity

$(v_{\hat{k},i})_i$ is balanced for every key hypothesis \hat{k} .

$\Rightarrow \hat{\sigma}(V_{\hat{k}})$ and $\hat{\mathbb{E}}(V_{\hat{k}})$ are constant w.r.t. \hat{k} and are perfect!



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SCA-reduction

A (\mathbf{m}, Δ) -SCA is **SCA-reducible** to a (\mathbf{m}', Δ') -SCA if there exists a function \mathbf{f} s.t. $\mathbf{m} = \mathbf{f} \circ \mathbf{m}'$ and for every (k, \hat{k}) and every samples $(\ell_{k,i})_i$ and $(v_{\hat{k},i})_i$, there exists a strictly monotonous function \mathbf{g} s.t.:

$$\Delta \left((\ell_{k,i})_i, (m_{\hat{k},i})_i \right) = \mathbf{g} \circ \Delta' \left((\ell_{k,i})_i, (m'_{\hat{k},i})_i \right) ,$$

where $m_{\hat{k},i} = \mathbf{m}(v_{\hat{k},i})$ and $m'_{\hat{k},i} = \mathbf{m}'(v_{\hat{k},i})$.

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single-bit DPA

- $\text{Im}(\mathbf{m}) = \{w_0, w_1\}$
- $\text{SB-DPA}(\hat{k}) = \hat{\mathbb{E}}(L \mid M_{\hat{k}} = w_0) - \hat{\mathbb{E}}(L \mid M_{\hat{k}} = w_1) .$



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Choose $\mathbf{f} = \text{Id}$ and $\mathbf{g} = \text{Id}$. In fact, just a rewriting!

PPA reduction

$$\text{SB-DPA}(\hat{k}) = \alpha_0 \hat{\mathbb{E}}(L \mid M_{\hat{k}} = w_0) + \alpha_1 \hat{\mathbb{E}}(L \mid M_{\hat{k}} = w_1) = \text{PPA}_{(\alpha_0, \alpha_1)}(\hat{k})$$

all-or-nothing DPA

- $\text{Im}(\mathbf{m}) = \{\dots, \omega_0, \dots, \omega_1, \dots\}$
- $\text{AON-DPA}(\hat{k}) = \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_0) - \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_1) .$



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PPA reduction

- $\alpha_0 = 1, \alpha_1 = -1$ and $\alpha_i = 0$ for every $\omega_i \in \text{Im}(\mathbf{m}) - \{\omega_0, \omega_1\}$
- $\text{AON-DPA}(\hat{k}) = \alpha_0 \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_0) + \alpha_1 \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_1) + \sum_{i \neq 0,1} 0 \cdot \alpha_i \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_i) = \text{PPA}_{(\alpha_i)_i}(\hat{k})$

generalized DPA

- $\text{Im}(m) = \dots \cup \Omega_0 \cup \dots \cup \Omega_1 \cup \dots$
- $\mathbf{G}\text{-DPA}(\hat{k}) = \hat{\mathbb{E}}(L \mid M_{\hat{k}} \in \Omega_0) - \hat{\mathbb{E}}(L \mid M_{\hat{k}} \in \Omega_1) .$



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PPA reduction

$$\begin{aligned}
 \mathbf{G}\text{-DPA}(\hat{k}) = & \\
 \sum_{\omega \in \Omega_0} \frac{\widehat{\text{Pr}}(M_{\hat{k}} = \omega)}{\widehat{\text{Pr}}(M_{\hat{k}} \in \Omega_0)} \widehat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega) & - \sum_{\omega \in \Omega_1} \frac{\widehat{\text{Pr}}(M_{\hat{k}} = \omega)}{\widehat{\text{Pr}}(M_{\hat{k}} \in \Omega_1)} \widehat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega) \\
 + \sum_{\omega \in \text{Im}(\mathbf{m}) \setminus \Omega_0 \cup \Omega_1} 0 \cdot \widehat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega) & = \mathbf{PPA}_{(\alpha_i)_i}(\hat{k})
 \end{aligned}$$

What about the PPA distinguisher?

PPA

- $\text{Im}(\mathbf{m}) = \{\omega_i; i\}$.
- $\text{PPA}_{(\alpha_i)_i}(\hat{k}) = \sum_{\omega_i \in \text{Im}(\mathbf{m})} \alpha_i \cdot \hat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_i)$.



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- After defining $\mathbf{f}(\omega_i) = \frac{\alpha_i}{\widehat{\mathbb{P}}_r(M_{\hat{k}} = \omega_i)}$ and $M'_{\hat{k}} = \mathbf{f}(M_{\hat{k}})$ we get:

$$\begin{aligned}
 \text{PPA}_{(\alpha_i)_i}(\hat{k}) &= \sum_{\omega_i \in \text{Im}(\mathbf{m})} \mathbf{f}(\omega_i) \cdot \widehat{\mathbb{P}}_r(M_{\hat{k}} = \omega_i) \cdot \widehat{\mathbb{E}}(L \mid M_{\hat{k}} = \omega_i) \\
 &= \sum_{\alpha \in \text{Im}(\mathbf{f})} \alpha \cdot \widehat{\mathbb{P}}_r(M_{\hat{k}} \in \mathbf{f}^{-1}(\alpha)) \cdot \widehat{\mathbb{E}}(L \mid M_{\hat{k}} \in \mathbf{f}^{-1}(\alpha)) \\
 &= \sum_{\alpha \in \text{Im}(\mathbf{f})} \widehat{\mathbb{P}}_r(M'_{\hat{k}} = \alpha) \cdot \widehat{\mathbb{E}}(\alpha \cdot L \mid M'_{\hat{k}} = \alpha) = \widehat{\mathbb{E}}(LM'_{\hat{k}})
 \end{aligned}$$

On the other side, the CPA distinguisher for the same model function $\mathbf{m}' = \mathbf{f} \circ \mathbf{m}$ gives:

$$\begin{aligned}
 \text{CPA}(\hat{k}) &= \frac{\hat{\mathbb{E}}(LM'_{\hat{k}}) - \hat{\mathbb{E}}(L)\hat{\mathbb{E}}(M'_{\hat{k}})}{\hat{\sigma}(L)\hat{\sigma}(M'_{\hat{k}})} \\
 &= \frac{1}{\hat{\sigma}(L)\hat{\sigma}(M'_{\hat{k}})} \cdot \hat{\mathbb{E}}(LM'_{\hat{k}}) - \frac{\hat{\mathbb{E}}(L)\hat{\mathbb{E}}(M'_{\hat{k}})}{\hat{\sigma}(L)\hat{\sigma}(M'_{\hat{k}})} \\
 &= a \cdot \hat{\mathbb{E}}(LM'_{\hat{k}}) + b
 \end{aligned}$$



CPA and PPA are SCA-equivalent

From

$$\text{PPA}_{(\alpha_i)_i}(\hat{k}) = \hat{\mathbb{E}} \left(LM'_{\hat{k}} \right)$$

and

$$\text{CPA}(\hat{k}) = a \cdot \hat{\mathbb{E}} \left(LM'_{\hat{k}} \right) + b$$

we deduce that the two attacks are equivalent thanks to a change of model $\mathbf{m} \rightarrow \mathbf{m}' = \mathbf{f} \circ \mathbf{m}$.



■ MIA

- ▶ Requires non-bijective function as sensitive variable
- ▶ Very expensive computational cost (pdf estimation)
- ▶ Less efficient than correlation-based attacks



- MIA
 - ▶ Requires non-bijective function as sensitive variable
 - ▶ Very expensive computational cost (pdf estimation)
 - ▶ Less efficient than correlation-based attacks
- Linear regression



Input

- A set of N input data $(x_i)_i$
 - ▶ for us $(p_i)_i$
- A set of N observations
 - ▶ in our case $(\ell_{k,i})_i$
- A set of $n < N$ basis functions $(g_j)_j$
 - ▶ For instance: $n = 9$, $g_0(x) = 1$ and $g_j(p_i) = v_{k,i}[j]$ the j^{th} bit of $v_{k,i}$



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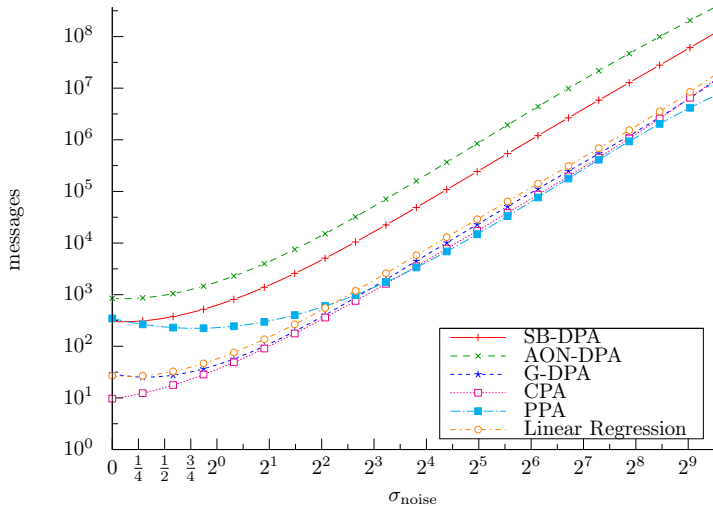
Output

- n values $(\hat{\beta}_i)_i \in \mathbb{R}$ that minimizes the Euclidean Distance between the vector $(\sum_j \hat{\beta}_j g_j(x_i))_i$ and the vector $(\ell_{k,i})_i$ i.e.

$$(\hat{\beta}_i)_i = \min_{(\beta_i)_i} \sum_i (\ell_{k,i} - \sum_j \beta_j \cdot v_{k,i})^2 .$$

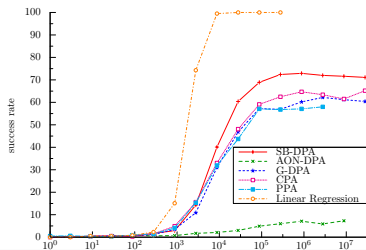
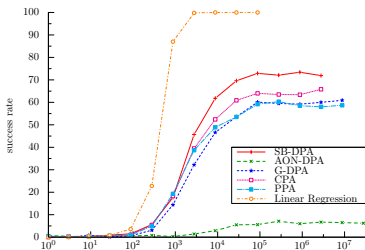
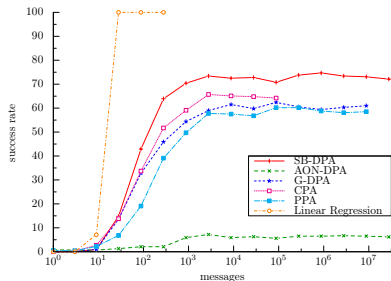


Perfect Model Scenario



Random Linear Leakage Scenario

$$\delta(x) = \varepsilon_0 + \sum_i^n \varepsilon_i x_i$$



- 1 Introduction
 - Problematics
 - Motivations
- 2 Description
 - Preliminaries
 - Description
- 3 Our Work - Analyses
 - All DPAs reduce to PPA
 - From PPA to CPA
 - Model-Free
 - Experiments
- 4 Conclusion



■ Linear Correlation

- ▶ Only One distinguisher is statistically sound: CPA
- ▶ “Enhance” CPA equals to refine the model
- ▶ It’s a matter of model



■ Linear Correlation

- ▶ Only One distinguisher is statistically sound: CPA
- ▶ “Enhance” CPA equals to refine the model
- ▶ It’s a matter of model

■ Linear Estimation

- ▶ Exhibit the model is the good way
- ▶ MIA is the *theoretically perfect* distinguisher
- ▶ Linear regression seems to be the *pratically perfect* distinguisher



- What about non linear leakages
 - ▶ Soundness in real life ?
 - ▶ Behavior of former distinguishers.
 - ▶ New statistical tools (e.g. non linear regression).
- Multivariate attacks
- A way to template attacks ?



That's all Folks!

